

Containment efficiency and control strategies for the Corona pandemic costs

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The rapid spread of the Coronavirus (COVID-19)¹ confronts policy makers with the problem of measuring the effectiveness of containment strategies and the need to balance public health considerations with the economic costs of a persistent lockdown. We introduce a modified epidemic model, the controlled-SIR model, in which the disease reproduction rate evolves dynamically in response to political and societal reactions. An analytic solution is presented. The model reproduces official COVID-19 cases counts of a large number of regions and countries that surpassed the peak of the outbreak. A single unbiased feedback parameter is extracted from field data and used to formulate an index that measures the efficiency of containment policies (the CEI index). CEI values for a range of countries are given. For two variants of the controlled-SIR model, detailed estimates of the total medical and socio-economic costs are evaluated over the entire course of the epidemic. Costs comprise medical care cost, the economic cost of social distancing, as well as the economic value of lives saved. Under plausible parameters, strict measures fare better than a hands-off policy. Strategies based on actual case numbers lead to substantially higher total costs than strategies based on the overall history of the epidemic.

In March 2020 the World Health Organization (WHO) declared the Coronavirus (COVID-19) outbreak a pandemic¹. In response to the growth of infections and in particular to the exponential increase in deaths², a large number of countries have been put under lockdown, with a considerable and potentially far reaching toll on economic activities³. In this situation it is paramount to provide scientists, the general public and policy makers with reliable estimates of both the efficiency of containment measures and the overall costs resulting from alternative strategies.

The societal and political response to a major outbreak like COVID-19 is highly dynamic, changing often rapidly with increasing case numbers. We propose to model the feedback of spontaneous societal and political reactions by a standard epidemic model that is modified in one key point: the reproduction rate of the virus is not constant, but evolves over time alongside with the disease in a way that leads to a ‘flattening of the curve’⁴. The basis of our investigation is the SIR (Susceptible, Infected, Recovered) model, which describes the evolution of a contagious disease for which immunity is substantially longer than the time-scale of the outbreak⁵. A negative feedback-loop between the severity of the outbreak and the reproduction factor g_0 is then introduced. As a function of the control strength α , which unites the effect of individual, social and political reactions to disease spreading, either an uncontrolled epidemic ($\alpha = 0$), or a strongly contained outbreak is described, as illustrated in Fig. 1a. The model is validated using publicly available COVID-19 case counts from an extended range of countries and regions. We provide evidence for data collapse when the case counts of distinct outbreaks are rescaled with regard to their peak values. A comprehensive theoretical description based on an analytic solution

of the controlled-SIR model is given. One finds substantial differences in the country-specific intrinsic reproduction factor and its doubling time. The controlled-SIR model allows in addition to formulate an unbiased benchmark for the effectiveness of containment measures, the containment efficiency index (CEI).

The controlled-SIR model is thoroughly embedded in epidemiology modeling. Early on, the study of the dynamics of measles epidemics⁶ has shown that human behavior needs to be taken into account^{7,8}. In this regard, a range of extensions to the underlying SIR model have been proposed in the past, such as including the effect of vaccination, contact-frequency reduction and quarantine⁹, human mobility¹⁰, self-isolation¹¹, the effects of social and geographic networks¹², and the influence of explicit feedback loops¹³. For an in-depth description, epidemiology models need to cover a range of aspects¹⁴, which prevents in general the possibility of an explicit analytic handling.

Political containment efforts, such as social-distancing measures, are implicitly included in the controlled-SIR model. We therefore use this model to estimate the overall economic and health-related costs associated with distinct containment strategies, accumulated over the entire course of an epidemic. This approach extends classical studies of the economic aspects of controlling contagious diseases. A central question is here the weighting of the economic costs of a disease transmission against the cost of treatment, and the loss of life, where a framework has been established^{15,16}. For the value of life, the statistical approach attributes the monetary value of an avoided premature death^{17–19}. This framework has been applied to the Corona pandemic in several recent contributions in which the evolution of the epidemic has generally been taken as exogenous²⁰ relying on estimates for the infec-

tion²¹, and case fatality rates^{22,23}. In this context, further studies have also discussed the relative effectiveness of control measures^{21,24}, and the possible future course of the disease²⁵.

Controlled-SIR Model

At a given time t we denote with $S = S(t)$ the fraction of susceptible (non-affected) individuals, with $I = I(t)$ the fraction of the population that is currently ill (active cases), and with $R = R(t)$ the fraction of recovered or deceased individuals. Normalization demands $S+I+R = 1$ at all times. The continuous-time SIR model²⁶

$$\tau \dot{S} = -gSI, \quad \tau \dot{I} = (gS - 1)I, \quad \tau \dot{R} = I \quad (1)$$

describes an isolated epidemic outbreak characterized by a timescale τ and a dimensionless reproduction factor g . Social and political reactions reduce the reproduction factor below its intrinsic (medical disease-growth) value, g_0 . This functionality can be described by

$$g = \frac{g_0}{1 + \alpha X}, \quad X = 1 - S. \quad (2)$$

The reaction to the epidemic is assumed to be triggered by the total fractional case count X (i.e. the sum of active, recovered and deceased cases), with α encoding the reaction strength. We note that empirical data on the reaction of governments and the severity of the outbreak suggest that this relation holds²⁷. Further below we will examine in addition strategies that are based, as an alternative framework, on the fraction of actual active cases, I .

The inverse functionality in equation (2) captures the notion that it becomes progressively harder to reduce g when increasing social distancing. Reducing g only somewhat is comparatively easy, a suppression by several orders of magnitude requires, in contrast, a near to total lockdown. We denote equation (1) together with (2) the controlled-SIR model. Key to our investigation is the observation that one can integrate the controlled-SIR model analytically, as shown in the Methods section, to obtain the phase-space relation

$$I = \frac{\alpha + g_0}{g_0} X + \frac{1 + \alpha}{g_0} \log(1 - X). \quad (3)$$

This relation, which we denote the ‘XI representation’, is manifestly independent of the time scale τ .

The medical peak load I_{peak} of actual infected cases is reached at a total fractional case count $X = X_{\text{peak}}$, which is given by

$$gS = 1, \quad X_{\text{peak}} = \frac{g_0 - 1}{g_0 + \alpha}, \quad (4)$$

For $\alpha = 0$ (no control), X_{peak} reduces to the well-known result $X_{\text{peak}} = (g_0 - 1)/g_0$. I_{peak} is then obtained from equations (3) and (4),

$$I_{\text{peak}} = \frac{g_0 - 1}{g_0} + \frac{1 + \alpha}{g_0} \log\left(\frac{1 + \alpha}{g_0 + \alpha}\right). \quad (5)$$

For $\alpha = 0$, I_{peak} is sometimes called the ‘herd immunity point’. The XI representation can be parameterized consequently either by g_0 and α , as in equation (3), or indirectly by X_{peak} and I_{peak} , which are measurable. In Fig. 1a an illustration of the XI-representation is given. For $g_0 = 3$ (in units of the disease duration) and $\alpha = 0$ one has $X_{\text{peak}} = 2/3$ and $I_{\text{peak}} \approx 0.3$. The total fraction of infected X_{tot} is 94%, which implies that only about 6% of the population remains unaffected. Containment policies, $\alpha > 0$, reduce these values. Fig. 1a and equation (5) illustrate a sometimes encountered misconception regarding the meaning of the herd immunity point, which we have labelled simply I_{peak} . The epidemic doesn’t stop at I_{peak} since infections continue beyond this point, albeit at a declining rate.

XI representation of COVID-19 outbreaks

Regional COVID-19 outbreaks are described by the controlled-SIR model to an astonishing degree of accuracy. For the analysis presented in Fig. 1b,c we divided, as described in the Methods section, the official case counts by the nominal population size of the respective region or country. Five-day centered averages are performed in addition. The such obtained country- and region specific XI representations are then fitted by equation (3).

It has been widely discussed that official case counts are affected by a range of factors, which include the availability of testing facilities and the difficulty to estimate the relative fraction of unreported cases^{28,29}. For example, as of mid-March 2020, the degree of testing for COVID-19, as measured by the proportion of the entire population, varied by a factor of 20 between the United States (340 tests per million) and South Korea (6100 tests per million)³⁰. The true incidence might be, according to some estimates³¹ higher by up-to a factor of ten than the numbers reported in the official statistics as positive. Since in the XI representation the number cases per population enters, scaling both I and X with a constant undercounting factor can be absorbed by an appropriate compensating renormalization of α , which is implicit in the fits shown in Fig. 1b,c. The XI framework is in this sense robust. Renormalization becomes however invalid if the undercounting of infection cases changes abruptly at a certain point during the epidemics, f.i. as a result of substantially increased testing. We will come back to this point further below. A fundamental change in the strategy followed by the government, e.g. from laissez faire to restrictive, would lead likewise to a change in α , which is not captured in the current framework.

In Table I we present for a number of countries and regions the obtained native dimensionless growth factors g_0 and the corresponding doubling times τ_2 , where $\tau_2 = \log(2)/\log(g_0)$ defines the number of time units τ needed to double case numbers.

For a robustness check we evaluated the parameters of the controlled-SIR model assuming that only a fraction f of the nominal population of the country or region

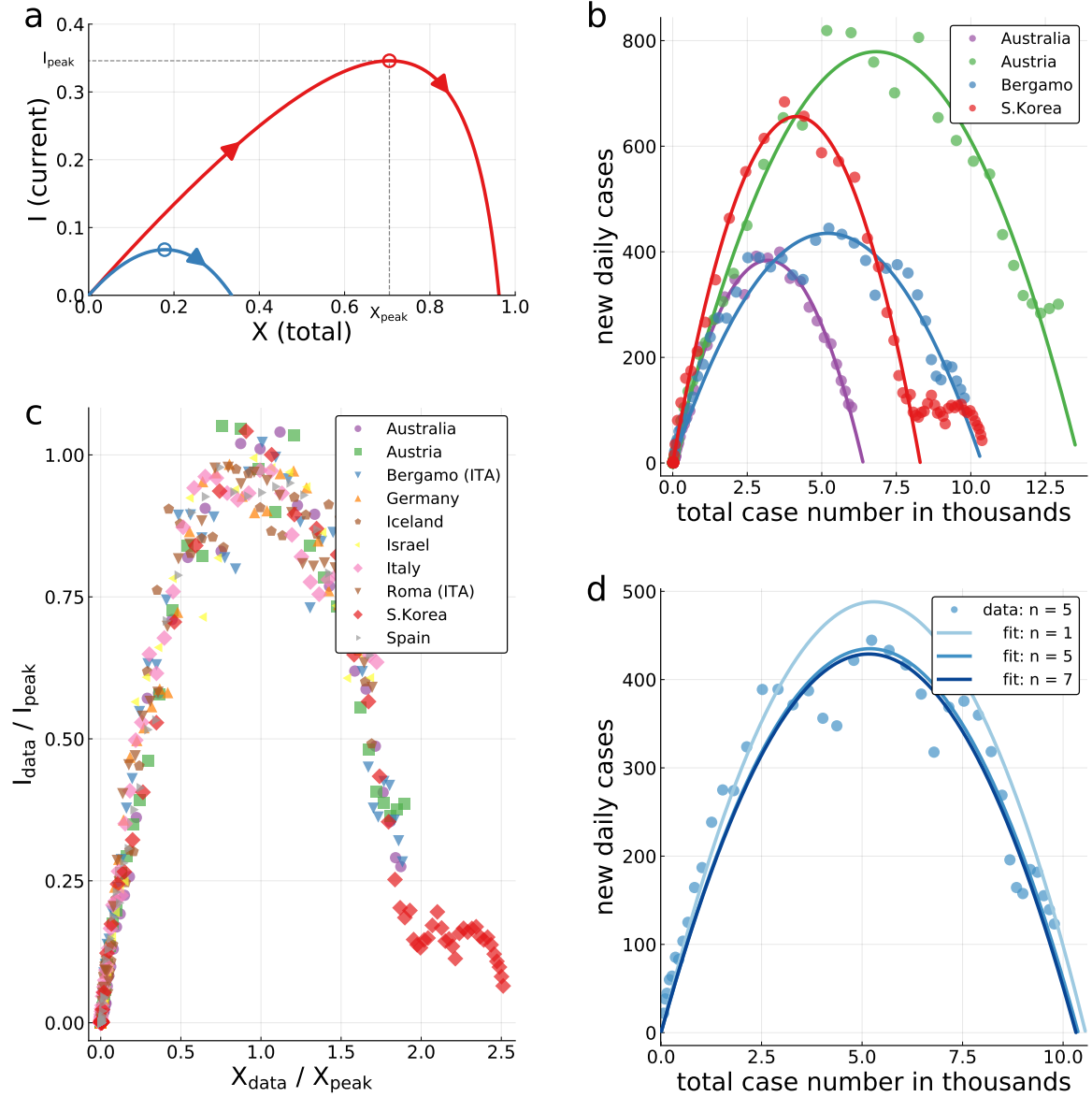


FIG. 1. **XI representation of COVID-19 outbreaks.** **a**, Model illustration. The closed expression $I = I(X)$ of actual infected cases I as a function of total infected cases X , as given by equation (3), is shown for two cases: $\alpha = 0$ (no control), red line and $\alpha = 10$ (long-term control), blue line for an intrinsic reproduction factor (in units of the disease duration) of $g_0 = 3$. The number of infections is maximal at I_{peak} (open circle), after starting at $X = I = 0$, with the epidemic ending when the number of actual cases drops again to zero. At this point the number of infected reaches X_{tot} . The peak of the uncontrolled case ($\alpha = 0$), $X_{\text{peak}} = 2/3$, is sometimes called the ‘herd immunity’ point. The final fraction of infected is $X_{\text{tot}} = 0.94$. **b**, Model validation for a choice of four countries/regions. The model (lines) fits the five-day centered averages of COVID-19 case counts well. For South Korea data till March 10 (2020) has been used for the XI-fit, at which point a transition from long-term overall control to the tracking of individuals is observable. **c**, Data collapse shown for ten countries/regions. Rescaling with the peak values X_{peak} and I_{peak} , obtained from the XI fit, maps COVID-19 case counts approximately onto a universal inverted parabola. **d**, Robustness test. The often strong daily fluctuations are smoothed by n -day centered averages. Shown are the Bergamo data (dots, $n = 5$) and XI-fits to $n = 1$ (no average), $n = 5$ and $n = 7$. Convergence of the XI-representation is observed.

in question could be potentially infected, possibly due to the presence of social or geographical barriers to the disease spreading. Only marginal differences were found for $f = 1/3$. The data presented in Table I suggest that there is a substantial spread in the country-specific intrinsic doubling times τ_2 . A direct connection between the severeness of an outbreak and the length of the respective doubling time is not evident. Regarding the US, which did not yet pass the peak for the most part, the analysis is preliminary. For New York City, an early estimate of the CEI value indicates a value of approximately 0.96, below the one of Bergamo.

Data collapse for COVID-19

Given that the XI representation is determined solely by two quantities, X_{peak} and I_{peak} , universal data collapse can be attained by plotting field data normalized with regard to the respective peak values, viz by plotting I/I_{peak} as a function of X/X_{peak} . It is remarkable, to which degree the country- and region specific official case counts coincide in relative units, see Fig. 1c. It implies that the controlled-SIR model constitutes a faithful phase-space representation of epidemic spreading subject to socio-political containment efforts.

Asymmetry of up-/down time scales

For the controlled SIR model an explicit analytic expression for the $X - I$ phase space representation can be derived, as given by equation (3), but not for the complete timeline $X(t)$ and $I(t)$. Exploiting the fact that fractional case counts are small for real-world epidemic outbreaks, the universal relation

$$\frac{\text{time down from the peak}}{\text{time up to the peak}} = 2g_0 - 1 \quad (6)$$

TABLE I. **COVID-19 containment efficiency index.** For selected countries, key COVID-19 parameters, as extracted from the respective official case counts. Given is the intrinsic dimensionless growth factor g_0 , the doubling time $\tau_2 = \log(2)/\log(g_0)$, in units of τ , and the containment efficiency index $\text{CEI} = \alpha/(g_0 + \alpha)$.

| location | | g_0 | τ_2 | CEI |
|---------------|-----|-------|----------|-------|
| Italy | ITA | 1.17 | 4.4 | 0.991 |
| Iceland | ISL | 1.19 | 4.0 | 0.983 |
| Bergamo | ITA | 1.20 | 3.8 | 0.972 |
| Roma | ITA | 1.20 | 3.8 | 0.998 |
| Germany | DEU | 1.21 | 3.6 | 0.995 |
| United States | USA | 1.22 | 3.5 | 0.994 |
| Spain | ESP | 1.23 | 3.3 | 0.990 |
| Luxembourg | LUX | 1.28 | 2.8 | 0.988 |
| Austria | AUT | 1.30 | 2.6 | 0.997 |
| Israel | ISR | 1.30 | 2.6 | 0.997 |
| Australia | AUS | 1.32 | 2.5 | 0.999 |
| South Korea | KOR | 1.46 | 1.8 | 1.000 |

between the time the outbreak needs to retreat from the peak, and to reach it in first place, can however be found, as shown in the Methods section. Interestingly, the ratio of down-/ and up-times is independent of the control strength α , and hence valid for epidemic outbreaks in general. For COVID-19, typical values of g_0 are of the order of 1.2-1.3, as listed in Table I, which implies that outbreaks take of the order of 40-60% longer to retreat than to ramp up.

Containment efficiency index

The control strength α enters the reproduction factor as αX , see equation (2). Data collapse suggest that regional and country-wise data is comparable on a relative basis. From $\alpha X = (\alpha X_{\text{peak}})(X/X_{\text{peak}})$ it follows that $\alpha X = \alpha(g_0 - 1)/(g_0 + \alpha)$ is a quantity that measures the combined efficiency of socio-political efforts to contain an outbreak. Dividing by $g_0 - 1$ results in a normalized index, the ‘Containment Efficiency Index’ (CEI):

$$\text{CEI} = \frac{\alpha X_{\text{peak}}}{g_0 - 1} = \frac{\alpha}{g_0 + \alpha}, \quad (7)$$

with $\text{CEI} \in [0, 1]$. The index is unbiased, being based solely on case count statistics, and not on additional socio-political quantifiers. Our estimates are given in Table I. The values for the evaluated regions/ countries are consistently high, close to unity, the upper bound, indicating that the near-to-total lockdown policies implemented by most countries has been effective in containing the spread of COVID-19. A somewhat reduced CEI value is found for the particularly strongly affected Italian region of Bergamo. For South Korea the CEI is so high that its deviation from unity cannot be measured with confidence.

Long-term vs. short-term control

So far, in equation (2) it was assumed that society and policy makers react to the total case count of infected X . This reaction pattern, which one may denote as ‘long-term control’, describes field data well. It is nevertheless of interest to examine an alternative, short-term control:

$$g = \begin{cases} g_0/(1 + \alpha I) & \text{(short-term)} \\ g_0/(1 + \alpha X) & \text{(long-term)} \end{cases} \quad (8)$$

For short-term control the relevant yardstick is given by the actual case number of infected I . The SIR model with short-term control cannot be analytically integrated, in contrast to the long-term controlled SIR model. It is therefore convenient to use the discrete-time controlled-SIR model,

$$I_{t+1} = \rho_t I_t (1 - X_t), \quad X_t = \sum_{k=0}^{\infty} I_{t-k}, \quad (9)$$

for numerical simulations. The reproduction factor ρ_t is given by the discrete-time version of equation (8), with

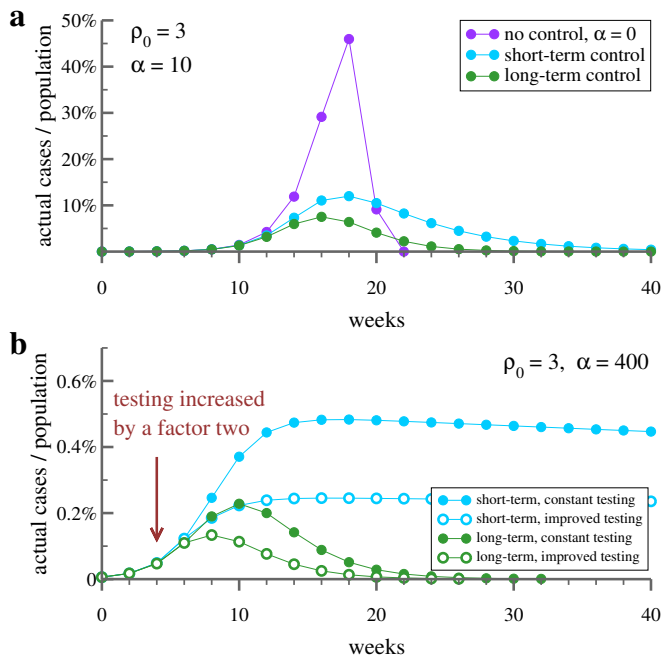


FIG. 2. **Control of epidemic peak.** **a**, Shown is the timeline of actual infected cases during an epidemic outbreak with an intrinsic reproduction factor (in units of the disease duration) of $\rho_0 = 3.0$, which is close to COVID-19 estimates³². The simulation is obtained by iterating equation (9), with one iteration corresponding to two weeks, taken as the average duration of the illness. Short-term control, which responds to the actual number of cases, see equation (8), is able to reduce the peak strain on the hospital system, but only by prolonging substantially the overall duration. Long-term control, which takes the entire history of the outbreak into account, is able to reduce both the peak and the duration of the epidemic. **b**, Increasing testing by a factor two (arrow), reduces the undercounting factor which increases, in turn, the effective response strength for both, the peak number of actual cases and the duration of the outbreak.

a ρ_0 that is appropriate for the time period associated with one transition in equation (9). Fig. 2 illustrates the capability of short-term and long-term reaction policies to contain an epidemic. While both strategies are able to lower the peak of the outbreak with respect to the uncontrolled ($\alpha = 0$) case, the disease will become close to endemic when the reaction is based on the actual number of cases, I_t , and not on the overall history of the outbreak.

Also included in the lower panel of Fig. 2 is a protocol simulating an increase of testing by a factor of two. Here $\alpha = 400$ has been used as the starting reaction strength, which increases by a factor of two when testing reduces the undercounting ratio by one half. One observes that long-term control is robust, in the sense that increased testing contributes proportionally to the containment of the outbreak. Strategies reacting to daily case number are in contrast likely to produce an endemic state.

The framework developed here, equations (1) and (2), describes mass control strategies, which are necessary

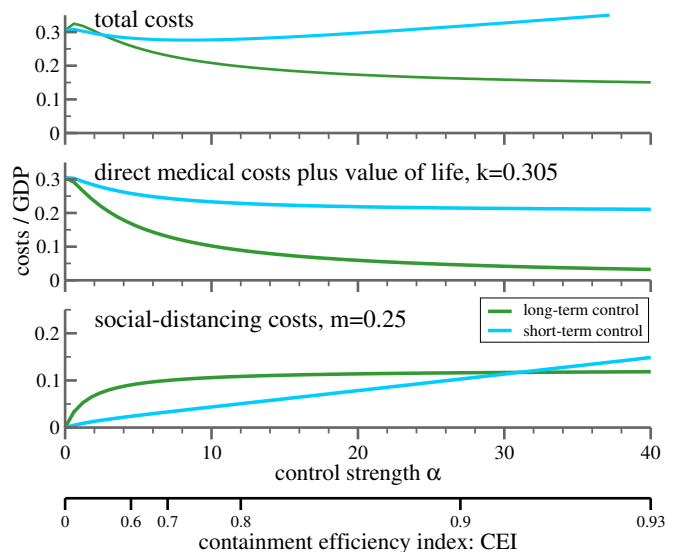


FIG. 3. **Cost of epidemic control strategies.** Shown are the costs in terms of $\text{GDP}_{\text{p.c.}}$ for long-term and short-term control, as defined by equation (8), both as a function of α and the CEI values (7), as indicated by the additional axis at the bottom. Given are the costs incurring from social distancing, equation (10) with $m = 0.25$ (lower panel), the pure medical costs with value of life costs (middle panel) and the sum of social and medical costs (upper panel). It is assumed that the containment policy switches from mass control to individual tracking when the fraction of actual cases I_t drops below a threshold of $I_{\min} = 10^{-5}$. The starting $I_0 = 2 \cdot 10^{-5}$.

when overly large case number do not allow to track individual infections. The framework is not applicable once infection rates are reduced to controllable levels by social distancing measures. The horizontal ‘tail’ evident in the data from South Korea in Fig. 1b can be taken as evidence of such a shift from long-term mass control to the tracking of individual cases.

Costs of controlling the COVID-19 pandemic

As shown above, the controlled-SIR model allows for a faithful modeling of the entire course of an outbreak. We apply it now to investigate how distinct policies and societal reaction patterns, as embedded in the parameter α , influence the overall costs of the epidemic. This is an inter-temporal approach since the cost of restrictions today to public life (lockdowns, closure of schools, etc.) must be set against future gains in terms of lower infections (less intensive hospital care, fewer deaths). Four elements dominate the cost structure: (i) The working time lost due to an infection, (ii) the direct medical costs of infections, (iii) the value of life costs, and (iv) the cost related to ‘social distancing’. The first three are medical or health-related. All costs can be scaled in terms of GDP per capita ($\text{GDP}_{\text{p.c.}}$). This makes our analysis applicable not only to the US, but to most countries with similar $\text{GDP}_{\text{p.c.}}$, e.g. most OECD countries.

Overall cost estimates

The costs estimates, which are given in detail in the Supplementary Information, can be performed disregarding discounting. With market interest rates close to zero and the comparatively short time period over which the epidemic plays out, a social discount rate between 3% and 5% would make little difference over the course of one year³³.

Total health costs C^{medical} incurring over the duration of the epidemic are proportional to the overall fraction $X_{\text{tot}} = X_{t \rightarrow \infty}$ of infected, with a factor of proportionality k . We hence have $C^{\text{medical}} = kX_{\text{tot}}$. We estimate $k \approx 0.305$ in terms of $\text{GDP}_{\text{p.c.}}$ when all three contributions (working-time lost, direct medical cost, value of life) are taken into account, and $k \approx 0.14$ when value of life costs are omitted.

The economic costs induced by social-distancing measures, C^{social} , depend in a non-linear way on the evolution of new cases (short-term control) or the percentage of the population infected (long-term control). To be specific, we posit that the reduction of economic activity is percentage-wise directly proportional to the reduction in the reproduction factor, viz to $(1 - \rho_t/\rho_0)$:

$$C^{\text{social}} = \sum_{I_t > I_{\text{min}}} c_t^s, \quad c_t^s = m \left[1 - \frac{\rho_t}{\rho_0} \right] \frac{2}{52}, \quad (10)$$

where $2/52$ is the per year fraction of 2-week quarantine period, and the epidemic is considered to have stopped when the fraction of new infections I_t falls below a minimal value I_{min} . As detailed out in the Supplementary Information, a comprehensive analysis yields $m \approx 0.3$ in terms of $\text{GDP}_{\text{p.c.}}$. Note that the ansatz equation (10) holds only when mass control is operative, viz when large case numbers do not allow the tracking of individual infections.

Once k and m are known, one can compare the total costs incurring as the result of distinct polices by computing the sum of future costs for different values for α in equation (2). This is illustrated in Fig. 3 with the value of life costs included ($k = 0.305$), and in Extended Data Fig. 3, without value of life costs ($k = 0.14$). Given are the total cumulative costs for the two strategies considered, long-term and short-term control, both as a function of the respective implementation strength, as expressed by the value of α .

The middle panel of Fig. 3 shows that a society focused on short-term successes will incur substantially higher medical costs, because restrictions are relaxed soon after the peak. By contrast, if policy (and individual behavior) is influenced by the total number of all cases experienced so far, restrictions will not be relaxed prematurely and the medical costs will be lower for all values of α . The bottom panel shows the social distancing costs as a fraction of $\text{GDP}_{\text{p.c.}}$, which represent a more complicated trade-off between the severity of the restrictions and the time they need to be maintained. If neither policy, nor individuals react to the spread of the disease ($\alpha = 0$) the epidemic will take its course and costs are solely medical. This changes as soon as society reacts, i.e. as α

increases. Social distancing costs increase initially (i.e. for small values of α), somewhat stronger for the long-term than for the short-term reaction framework. The situation reverses for higher values of α with $\alpha \approx 30$ being the turning point. From there on, the distancing cost from a long-term based reaction falls below that of the short-term strategy. The sum of the two costs is shown in the uppermost panel. For large values of α , short-term policies result in systematically higher costs.

Discussion

The total costs of competing containment strategies can be estimated if the feedback of socio-political measures can be modeled. For this program to be carried out one needs two ingredients, (i) a validated epidemiological model and (ii) a procedure relating the success of containment efforts, in terms of model parameters, to the economic costs generated by the measures. Regarding the first aspect, we studied the controlled-SIR model and showed that COVID-19 outbreaks follow the phase-space trajectory, denoted the XI representation, predicted by the analytic solution. The same holds for the 2015 MERS outbreak in South Korea, as shown in Extended Data Fig. 1b. We extracted for a number of countries and regions accurate estimates for the intrinsic doubling times and found that they are not correlated to the severity of the outbreak.

Regarding the second aspect, we proposed that a percentage-wise corresponding principle between the achieved reduction in the infection rate and economic slowdown, equation (10), is able to establish the required link between epidemiology, political actions and economic consequences. Health-related costs, which are related to official case counts, are in contrast comparatively easier to estimate.

A non-trivial outcome of our study is that strong suppression strategies lead to lower total costs than taking no action, when containment efforts are not relaxed with falling infection rates. A short-term control approach of softening containment with falling numbers of new cases is likely to lead to a prolonged endemic period. With regard to the ‘exit strategy’ discussion, these findings imply that social distancing provisions need to be replaced by measures with comparative containment power. A prime candidate is in this regard to ramp up testing capabilities to historically unprecedented levels, several orders of magnitude above pre-Corona levels. The epidemic can be contained when most new cases can be tracked, as implicitly expressed by the factor α . This strategy can be implemented once infection rates are reduced to controllable levels by social distancing measures. Containment would benefit if the social or physical separation of the ‘endangered’ part of the population from the ‘not endangered’ would be organized in addition on a country-wide level, as suggested by community-epidemiology. With this set of actions the vaccine-free period can be bridged.

As a last note, there is a sometimes voiced misconception regarding the meaning of the herd immunity point,

which occurs for an infection factor of three when 66% of the population is infected. Beyond the herd immunity point, the infected-case counts remain elevated for a considerable time. The outbreak stops completely only once

94% of the population has been infected, as illustrated in Fig. 1a. The view that the epidemic is essentially over once the herd immunity point is reached is erroneous.

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Methods

Data collection and handling

Data has been accessed April 12 (2020) via the public COVID-19 Github repository of the John Hopkins Center of Systems Science and Engineering³⁴. Preprocessing was kept minimal, comprising only a basic smoothing with sliding averages. If not stated otherwise, a five time centered average (two days before/after plus current day) has been used. Robustness checks with one, three and seven days were performed, as shown in Fig. 1d. Fractional case counts are obtained by dividing the raw number by the respective population size. For the case of South Korea, the XI-analysis was performed using the initial outbreak, up to March 10 (2020). China has been omitted in view of the change in case count methodology mid February 2020.

The variable I represents in the SIR model the fraction of the population that is infectious. For the COVID-19 data, we used instead an XI-representation for which the number of new daily cases is plotted against the total case count. This procedure is admissible as long as the relative duration of the infectious period does not change.

Fitting procedure

Field data is crowded at low levels of X and I in the XI representation. A fitting routine that takes the range $X \in [0, X_{\text{tot}}]$ uniformly into account is attained when minimizing the weighed loss function

$$U = \sum_t u_t \left(I_t^{(\text{data})} - I^{(\text{theory})}(X_t^{(\text{data})}) \right)^2. \quad (11)$$

For the weight we used $u_t = X_t^{(\text{data})} - X_{t-1}^{(\text{data})} = I_t^{(\text{data})}$, which satisfies the sum-rule $\sum_t u_t = X_{\text{tot}}$. With equation (11) it becomes irrelevant where the timeline of field data is truncated, both at the start or at the end. Adding a large number of null measurements after the epidemic stopped would not alter the result. Numerically the minimum of U as a function of g_0 and α is evaluated.

Modeling field data as uncontrolled outbreaks

It is of interest to examine to which degree official case statistics could be modeled using an uncontrolled model, $\alpha = 0$. For this purpose it is necessary to assume that the epidemics stops on its own, which implies that one needs to normalize the official case counts not with respect to the actual population, but with respect to a fictitious population size N . In this view the outbreak starts and ends in a socially or geographically restricted community. The results obtained when optimizing N is included in Extended Data Fig. 1a. On first sight, the $\alpha = 0$ curve tracks the field data. Note however the very small effective population sizes, which are found to be 36 thousands for the case of Austria. Alternative one may adjust g_0 by hand during the course of an epidemic, as it is often done when modeling field data.

Analytic solution of the controlled-SIR model

Starting with the expression for the long-term control, equation (2), one can integrate the controlled-SIR model equation (1) to obtain a functional relation between S and I . Integrating \dot{I}/\dot{S} , viz

$$dI = -dS + \frac{1}{g(S)S} dS = -dS + \frac{1}{g_0} \frac{1 + \alpha(1 - S)}{S} dS,$$

yields

$$I = - \left(\frac{\alpha}{g_0} + 1 \right) S + \frac{1 + \alpha}{g_0} \log(S) + c, \quad (12)$$

where the integration constant c is given by the condition $I(S = 1) = 0$. Substituting $S = 1 - X$ one obtains consequently the XI-representation equation (3). The analogous result for $\alpha = 0$ has been derived earlier³⁵. The number of actual cases, I , vanishes both when $X = 0$, the starting point of the outbreak, and when the epidemic stops. The overall number of cases, X_{tot} , is obtained consequently by the non-trivial root X_{tot} of equation (3), as illustrated in Fig. 1a. As a side remark, we mention that the XI representation allows us to reduce equation (1) to

$$\tau \dot{S} = - \frac{gS}{g_0} \left[(\alpha + g_0)(1 - S) + (1 + \alpha) \log(S) \right], \quad (13)$$

which is one dimensional. Integrating equation (13) with $g = g(S)$ yields $S = S(t)$, from which $I(t)$ follows via $\tau \dot{I} = (gS - 1)I$ and $R(t)$ from the normalization condition $S + I + R = 1$.

Large control limit of the XI representation

Expanding equation (3) in X , which becomes small when $\alpha \gg 1$, one obtains

$$I = \frac{1 + \alpha}{2g_0} X \left[2 \frac{g_0 - 1}{1 + \alpha} - X \right] + O(X^3), \quad (14)$$

which makes clear that the phase-space trajectory becomes an inverted parabola when infection fractions are small. As a consequence one finds

$$I \approx \frac{g_0 - 1}{g_0} X + O(X^2), \quad (15)$$

which shows that the slope $dI/dX = (g_0 - 1)/g_0$ at $X \rightarrow 0$ is independent of α and of the normalization procedure used for I and X . The first result was to be expected, as α incorporates the reaction to the outbreak, which implies that α contributes only to higher order. The dimensionless natural growth factor g_0 is hence uniquely determined, modulo the noise inherent in field data, by measuring the slope of the daily case numbers with respect to the cumulative case count.

From equation (14) one obtains

$$X_{\text{tot}}|_{\alpha \gg 1} \approx 2 \frac{g_0 - 1}{\alpha} \quad (16)$$

for the total number of infected X_{tot} in the large-control limit. In analogy one finds

$$I_{\text{peak}}|_{\alpha \gg 1} \approx \frac{(g_0 - 1)^2}{g_0 \alpha}, \quad X_{\text{tot}} \approx \frac{2g_0}{g_0 - 1} I_{\text{peak}} \quad (17)$$

from equation (3), and in comparison with equation (16).

Time scale asymmetry

From the one-dimensional representation (13) of the controlled SIR model one can estimate two characteristic time scales. For this one considers an initial relative infection status $f_X X_{\text{tot}}$, with $f_X > 0$ and $f_X \ll 1$.

- **run-up:** T_{up} , defined as the time needed to reach the peak starting from $X_{\text{start}} = f_X X_{\text{tot}}$.
- **run-down:** T_{down} , defined as the time needed to reach $X_{\text{end}} = (1 - f_X) X_{\text{tot}}$, down from the peak.

In general one needs to integrate equation (13) numerically. Given that real-world fractional case counts X are small, $X < X_{\text{tot}} \ll 1$, one can simplify (13), as for (14), obtaining

$$t - t_0 = \frac{\tau}{g_0 - 1} \log \left(\frac{X}{(X_{\text{tot}} - X)^{2g_0 - 1}} \right). \quad (18)$$

It follows directly that $T_{\text{down}}/T_{\text{up}} = 2g_0 - 1$. For a pathogen to spread its dimensional growth factor g_0 needs to be larger than unity, compare Table I. Going down takes hence substantially longer than ramping up.

Data availability

The COVID-19 data examined is publicly accessible via the COVID-19 Github repository of the John Hopkins Center of Systems Science and Engineering <https://github.com/CSSEGISandData/COVID-19>. Data for the 2015 MERS outbreak in South Korea is publicly available from the archive of the World Health organization (WHO), https://www.who.int/csr/disease/coronavirus_infections/archive-cases/en/.

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Additional information

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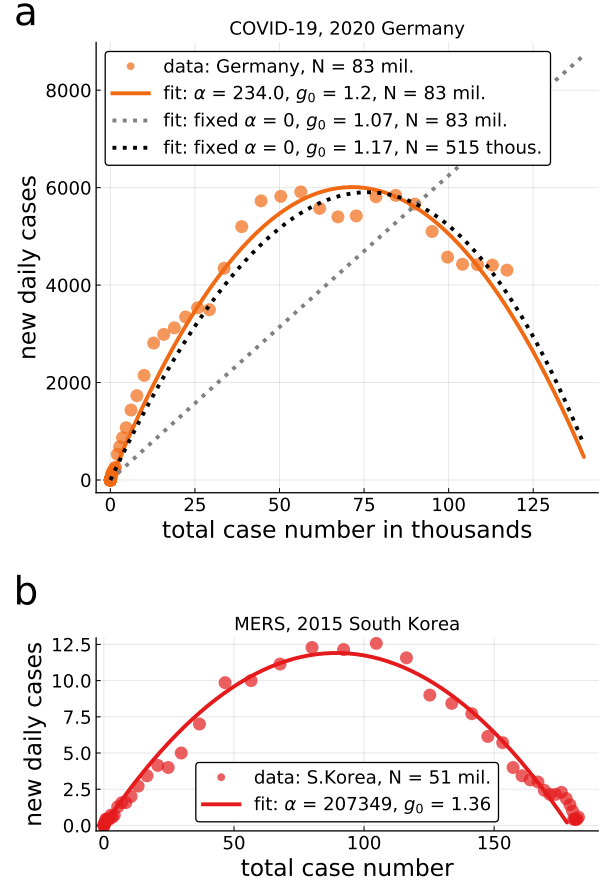


FIG. 4. **Extended Data Fig. 1** | **a**, Modeling case counts as uncontrolled outbreaks. Case counts, here for Germany (five-day centered averages, dots), can be modeled using either the full XI representation (full line), as given by equation (3), or with the standard uncontrolled SIR model ($\alpha = 0$, dashed lines). Using the nominal population size for Germany, 83 Million, leads to an utterly unrealistic $\alpha = 0$ curve (dashed, grey). The best $\alpha = 0$ fit is obtained when a fictitious population size of 478 Thousand is assumed (dashed, black). An epidemics abates on its own only when the population size is of the order of the total case count divided by X_{tot} . **b**, XI representation of the 2015 MERS outbreak in South Korea, covering a total of 186 cases. A $n = 7$ centered average has been used, in view of the small case numbers.

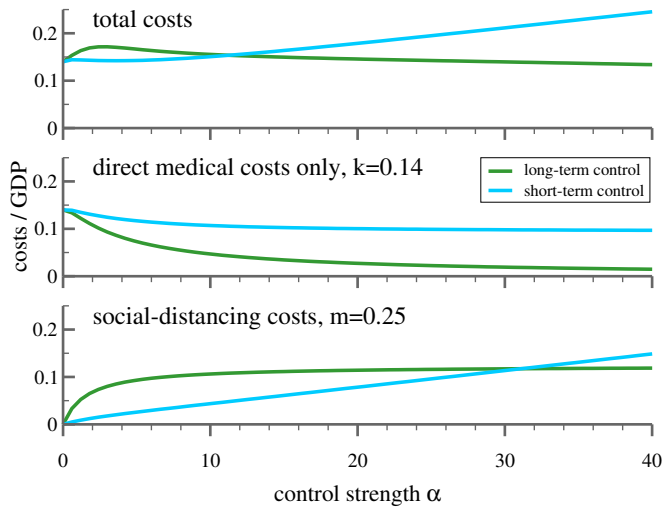


FIG. 5. **Extended Data Fig. 3 | Cost of epidemic control without value of life.** As in Fig. 3 (bottom panels are identical), but without the value of life costs. A long-term strategy with intermediate reaction strength is costlier than a hands-off policy.

SUPPLEMENTARY INFORMATION

Detailed costs of controlling the COVID-19 pandemic

In what follows we present a detailed estimation of the costs of controlling the COVID-19 pandemic given in GDP per capita ($GDP_{p.c.}$). We assume that four elements dominate the cost structure: (i) The working time lost due to an infection, (ii) the direct medical costs of infections, (iii) the value of life costs, and (iv) the cost related to ‘social distancing’. The first three are medical or health-related.

Health costs, loss of working time

A first direct impact of a wave of infections is that a fraction of the population cannot work. Based on the Diamond Princess data³⁶, where the entire population was tested, we estimate that only half of the infected develop symptoms that require them to stay home for a one- to two-week period and an additional two-week period until they are no longer contagious. About 20% of the total population (or 40% of those with symptoms) develop stronger symptoms requiring one additional period of absence from work²³. To be conservative, we assume that there are no severe cases or deaths among the working age. This results in a reduction in the work force per year (52 weeks) of around $(0.3 \times 2 + 0.2 \times 3) \times (2/52) = 2.4/52 = 5$ percent, for every 1% of the population infected.

Medical costs, treatment, hospitalization

There are no rigorous studies yet of the costs of treatment for the COVID-19, but it is estimated that about 20%³⁷ of the infected individuals require some sort of hospitalization, with around 5% needing intensive care and roughly 1% dying³⁶. As a comparison, we note that an average influenza season leads to an hospitalization of about 0.12% of the US population³⁸; and one fourth of them require intensive care, with one twentieth (0.13% of all infected) dying³⁶. Averaged over the 2010-17, of the order of 35 thousand influenza-related deaths per year have been registered in the US. For Germany, with a quarter of the US population, these numbers would translate into 8-9 thousands influenza deaths per year.

Intensive care with ventilation is the most costly form of life saving in hospital care. In the US, the cost of 2 weeks of an intensive care unit is equivalent to about 1 year (100%) of $GDP_{p.c.}$ ³⁹. In Germany, which might be typical of the rest of Europe, the cost of 2 weeks of intensive care appears to be somewhat lower, around 20,000 euro, or roughly 60% of $GDP_{p.c.}$ ⁴⁰. We use the German parameter for a conservative estimate of medical costs. The cost of general hospitalization for 2 weeks

is assumed to be 12,000 euro, and equivalent to about 30% of $GDP_{p.c.}$. It is estimated that the median time from onset to recovery for mild cases is approximately two weeks and 3-6 weeks for patients with severe or critical disease⁴¹. We use a conservative estimate of two weeks of intensive care and two weeks of general hospitalization for severe cases. This results in a medical cost of $(0.05 \times 0.6 + 0.05 \times 0.3 + 0.15 \times 0.3 = 0.09)$, that is 9% $GDP_{p.c.}$.

Value of lives lost

Third, the cost of premature death through the disease represents the most difficult contribution to evaluate in financial terms. We will show below that our central results remain valid even without assigning a value to lives lost, but since major contributions²⁰ are based on an evaluation of the economic value of lives lost, we show how this point can be incorporated into our framework. There are two ways to attribute a monetary value on a life saved or lost. The first one, mentioned above, is based on the concept of a Value of Statistical Life (VSL), which is commonly used in the impact assessment of public policy which aims at lowering the probability of an avoided premature death⁴². A typical application scenario for VSL is the case when the probability of death is very low (e.g. car accidents), but could be lowered even more (seat belts). For COVID-19, a high-death epidemic, we prefer a medical-based approach, which allows us to produce conservative estimates. VSL arrives in contrast often at much higher values, up to millions of euro or dollars⁴³. Putting a monetary value on lives saved is unavoidable in medical practice that is confronted with the problem of selecting the procedures to be used to prolong life - a situation that arises for many patients infected by the Coronavirus under intensive care. The literature dealing with the cost of medical procedures finds a central range of between 100,000 and 300,000 dollars per year of life lost^{44,45}. Given the current US $GDP_{p.c.}$, these values translate into a range of 1.5 to 4 years of $GDP_{p.c.}$. Cutler and Richardson⁴⁶ argue for a value equivalent to three times $GDP_{p.c.}$. We use the lower bound of this range for most of our simulations for a conservative estimate of the value of lives saved.

What remains to be determined is the number of years lost when a Corona patient dies. We rely on the data from the cruise ship Diamond Princess^{22,36} which served almost as a laboratory, the average age at death was 76 years. Cruise passengers tend to have fewer acute health conditions than the general population, thus rendering the co-morbidity argument less prominent. The remaining life expectancy (weighted by the difference incidence by sex) would thus be 11 years. This implies that the economic value of the premature deaths should be equal to about 11 times the loss for one year of life saved (potentially higher for most European countries which tend to have a higher life expectancy). For each 1% of the

population the value of lives lost would thus be equal to $0.01 \times 11 \times$ the nominal value of one year of life.

The value of life can be measured in terms of multiples of $\text{GDP}_{\text{p.c.}}$, which allows to write the sum of the three types of health or medical costs (loss of working time, hospitalization and value of lives lost) as a linear function of the percentage of the population infected:

$$c_t^{\text{med}} = kI_t$$

with a proportionality factor k being equal to the sum of the three contributions. Scaling k with the $\text{GDP}_{\text{p.c.}}$ allows for an application and comparison across countries. Using the lower bound of the central range yields then the following calibration of the medical costs:

$$(0.05 + 0.09 + 0.01 \times 1.5 \times 11) \times \text{GDP}_{\text{p.c.}} = 0.305 \times \text{GDP}_{\text{p.c.}} \quad (19)$$

The upper bound for the value of k would be substantially higher: $(0.05 + 0.09 + 0.01 \times 4 \times 11) \times \text{GDP}_{\text{p.c.}} = 0.58 \times \text{GDP}_{\text{p.c.}}$. For the numerical calculations we will use the conservative estimate $k = 0.305$ in terms of $\text{GDP}_{\text{p.c.}}$.

If we only consider the direct medical costs consisting of loss of working time and hospitalization, without including the value of lives lost, the proportionality factor in equation (19) reduces to $k = 0.14$ in terms of $\text{GDP}_{\text{p.c.}}$.

Medical costs over the lifetime of the epidemic

The cost estimates discussed so far, c_t^{med} , refer to the per-period cost of the currently infected. For the total cost over the entire epidemic we need to calculate the discounted sum of all c_t^{med} over time. Given that a period corresponds to about two weeks, we neglect discounting, which would make little difference even if one uses a social discount rate of 5% instead of using market rates (which may be negative). The total medical costs over the course of the endemic can be written as the simple sum of the cost per unit of time:

$$C^{\text{medical}} = \sum_{I_t > I_{\text{min}}} c_t^{\text{med}} = kX_{\text{tot}} \quad (20)$$

The epidemic is considered to have stopped when the fraction of new infections I_t falls below a minimal value, I_{min} .

Using the conservative estimate (low value of life) $k = 0.305$ it is straightforward to evaluate the total cost of a policy of not reacting at all to the spread of the disease, which would lead in the end to $X_{\text{tot}} = 0.94$. A hands-off policy would therefore lead to medical costs of over 28% of GDP.

In absolute terms the cost of a policy of doing nothing would amount to 1000 billion euro for a country like Germany. For the US the sum would be closer to 5 Trillion of dollars (25% of a GDP of 20 Trillion of dollars). As it would not be possible to ramp up hospital capacity in

the short time given the rapid spread of the disease, the cost would be in reality substantially higher, together with death toll^{20,21}. We abstract from the question of medical capacity (limited number of hospital beds) because we assume that society would react anyway as the virus spreads, thus limiting the peak, and, second, we are interested in the longer term implications of different strategies and not just in their impact on the short-term peak.

We note that even concentrating only on the direct medical cost and working time lost ($k=0.14$) a policy of letting the epidemic run its course through the entire population would lead to losses of working time and hospital treatment of over 13% of GDP (94% of 14%). By comparison, total health expenditure in most European countries amounts in normal times to about 11% of GDP⁴⁷. Even apart from ethical considerations, to avoid or not potentially hundreds of thousands of premature deaths, there exists thus an economic incentive to slow the spread of the COVID-19 virus.

Given the somewhat contentious nature of the value of lives lost, we present in the middle panel of Extended Data Fig. 3 of the main text the medical cost estimates (as a proportion of GDP) without including the value of life costs (results with including the value of life costs are shown in the main text). As shown in the figure, increasing α leads to a lower medical cost because the percentage of the population infected will be lower. The difference between short-term and long-term control increases for higher values of α . At these α values the medical cost over the entire endemic would be lower because the overall fraction of infected population is lower. For a strongly reactive society and policy i.e. for $\alpha \gg 1$ (and the case of long-term control), an explicit solution for the total health cost is given by,

$$C^{\text{medical}} = kX_{\text{tot}}|_{\alpha \gg 1} \approx 2k \frac{g_0 - 1}{\alpha} \quad (21)$$

which implies that the total health or medical costs are inversely proportional to the strength of the policy reaction parameter. Draconian measures from the start, i.e. α going towards infinity reduce the medical costs to close zero - irrespective of whether one adds the value of lives lost. This can be seen in Fig. 3 of the main text and the respective and Extended Data figure, where the medical cost (over the entire epidemic) starts for $\alpha = 0$ at values close to k because without any societal reaction 94% of the population would get infected and with increasing α the medical costs decline monotonously.

Social distancing costs

The economic costs of imposing social distancing on a wider population are at the core of policy discussions and drive financial markets. As mentioned above, social distancing can take many forms; ranging from abstaining

from travel or restaurant meals to government interventions enforcing lockdowns, quarantine, closure of schools, etc. This cost is more difficult to estimate. However, a rough estimate is possible if one takes into account that most economic activity involves some social interactions. Limiting social interaction thus necessarily reduces economic activity. This suggests that the economic cost of the social distancing described in equation (2) of the main text should increase with the reduction in the transmission rate described by g .

Without any social distancing, $\alpha = 0$, the economy would not be affected by the spread of the virus. Stopping all economic social interactions would bring the economy to a halt, but the reproduction rate of the virus would also go close to zero (Eichenbaum *et al.*⁴⁸ make a similar assumption). We thus posit that the (per-time unit) social-distancing economic cost c_t^s is proportional to the reduction in the transmission rate. The total economic costs C^{social} can be written as the sum of c_t^s :

$$C^{\text{social}} = \sum_{I_t > I_{\min}} c_t^s, \quad c_t^s = m \left[1 - \frac{\rho_t}{\rho_0} \right] \frac{2}{52} \quad (22)$$

considering here the notation of the discrete-time controlled SIR model (equation (8)). The key question is the factor of proportionality, m , which links the severity of social distancing to the reduction in economic activity. Popular attention has focused on services linked directly to social contact. There exist indeed selected sectors which will completely shut down under a lockdown. However, these sectors (tourism, non-food retail, etc.) account for a limited share of the economy (less than 10% for most countries). Expenditure for food is actually little affected since even under the most severe lockdown, grocery shopping is still allowed and families must consume more food at home as they cannot go out to restaurants.

The manufacturing sector is less affected by social distancing than the service sector because in modern factories workers are scattered over a large factory floor, making it relatively easy to maintain production while maintaining the appropriate distance between workers. Moreover, some sectors, e.g. finance, can work online with only a limited effect on productivity. The widespread impression that the entire economy stops under a lockdown is thus not correct. The drastic measures adopted in China illustrate this proposition: when all non-essential social interactions were forbidden, industrial production and retail sales fell by 'only' 20-25%⁴⁹ while the reproduction factor went from 3 to 0.3, a fall by a factor of ten. Using this experience we calibrate the parameter m at 0.25.

A reduction in the reproduction factor ρ_t to one tenth its normal epidemiological value of ρ_0 would thus lead to a loss of GDP of 25% for the time period during which the restriction or social distancing measures are in place. This would imply that an abrupt shutdown of the economy to 25% of its capacity for 12 weeks, or 6 incubation

periods would cost about $0.25 \times (12/52)$, or about 6% of annual GDP. A reduction of GDP by 6% would represent a recession even deeper than the one which followed the financial crisis of 2009. This is compatible with current forecasts of zero GDP growth in China in 2020 (relative to a baseline of 5-6% before the crisis). But even such a large cost in terms of output foregone would be below the medical cost arising from herd immunity. Even apart from ethical considerations, it would thus appear to make sense to accept a temporary shut down of parts of the economy to avoid the huge medical costs.

A first result is thus that if one compares two extremes: letting contagion run its course (herd immunity) or draconian measures, the social costs are lower in the second case. Small changes to the key parameters, k and m , might change the exact values of the costs in terms of overall magnitude, but the ranking appears robust.

We do not consider separately the fiscal cost, i.e. the cost for the government to save millions of enterprises from bankruptcy and ensure that workers have a replacement income when they get laid off. This cost to governments is a transfer within the country from one part of society (tax payers) to those who suffer most under the economic crisis.

A key issue in the discussion on the economic cost of social distancing is the question about how long these measures need to be maintained. It is sometimes argued that the cost of a policy of social distancing would be unacceptably high because the measures could not be relaxed until the virus had been totally eradicated. However, this pessimism is not warranted by the success of a strategy of 'testing and tracing' implemented in some countries (mainly those which had experienced SARS). Such a strategy is, of course, only possible if the starting number of infections is low enough to allow for individual tracing.

We thus make the assumption that when the number of active cases falls below a certain threshold, the costly measures of general social distance containment are no longer needed and can be substituted by pro-active repeated testing coupled to quick follow-up of the remaining few cases which are quarantined and whose contacts are quickly traced. In this case the resulting economic cost is assumed to fall away. The experience of Singapore and Japan suggests that when the infected are less than one per 100,000, general social distancing is no longer required (assuming mass testing has been adopted in the meantime so that the infections can be accurately measured).

Parameter updating

The estimates on which our results are based will have to be updated when actualized COVID-19 data is available in the future. The WHO-China Joint Mission Report suggests a ρ_0 (g_0 in the continuous-time representation) per infected of $2 - 2.5$ ⁴¹ (in units of the disease duration), while we use the figures from Liu *et al.*³², who predict a reproduction factor of around three. The numbers for

the forecast of health costs are derived in part from the Diamond Princess data³⁶, for which the population was comparatively healthy. The statistics for symptoms requiring the absence from work may therefore in reality be somewhat higher. The hospitalization and mortality rate are estimates with a substantial uncertainty, due to the high numbers of unregistered and untested infections. Early studies based on official data from China^{50,51} estimated that the number of actual infections may be between 10 to 20 times higher than the number of detected infections. However, first sample test in e.g. Austria suggest only a factor of 3⁵². Leaving possibly lower, but still substantial true hospitalization and mortality rates for COVID-19. One of our main goals has therefore been the introduction of a generic framework, which can be updated by future advances in the accuracy of estimates while still presenting specific results with the data available at this time.

Relation to further studies

A range of determining factors have been examined for the ongoing COVID-19 epidemic, in particular the effect of quarantine⁵³ and that community-level social distancing may be more important than the social distancing of individuals⁵⁴. An agent-based model for Australia found, in this regard, that school closures may not be decisive⁵⁵. Microsimulation models suggest, on the other hand, that a substantial range of non-pharmaceutical interventions are needed for an effective containment of the COVID-19 outbreak²¹.

We also note that there haven't been attempts to derive disease transmission rates from economic principles of behavior⁴⁸, which would allow to measure the cost of the Corona pandemic under different policy settings. Another strand of the literature takes the pandemic as given, and as the basis for scenarios for the economic impact and for the financial-market volatility⁵⁶⁻⁵⁸.