On the Sustainability of Currency Boards: Evidence from Argentina and Hong Kong
On the Sustainability of Currency Boards: Evidence from Argentina and Hong Kong\textsuperscript{a}

Chun-Yu Ho and Wai-Yip Alex Ho\textsuperscript{b}

Georgia Institute of Technology and Boston University

Version: September 2008

Abstract: This paper examines the sustainability of the currency board arrangements in Argentina and Hong Kong. We employ a Markov switching model with two regimes to infer the exchange rate pressure due to economic fundamentals and market expectations. The empirical results suggest that economic fundamentals and expectations are key determinants of a currency board’s sustainability. We also show that the government’s credibility played a more important role in Argentina than in Hong Kong. The trade surplus, real exchange rate and inflation rate were more important drivers of the sustainability of the Hong Kong currency board.

\textsuperscript{a} We thank Hans Genberg, Stefan Gerlach, Francois Gourio, Slavi T. Slavov, James Yetman, Matthew Yiu, Jeremy Smith and Michael Luca for helpful comments and discussion. We thank the Hong Kong Institute of Monetary Research for their hospitality during our visit in June 2008. We also thank the seminar participants at the Hong Kong Institute of Monetary Research, the Midwest Economics Association Annual Meeting 2007, Hong Kong’s Currency Board at 25 and the Southern Economics Association Annual Meeting 2008 for valuable comments.

\textsuperscript{b} Email: chunyu.ho@econ.gatech.edu for Chun-Yu Ho and hwylex@bu.edu for Alex Ho.
1 Introduction

Calvo and Reinhart (2002) find that a fixed exchange rate is a popular choice of exchange rate policy for developing economies. Maintaining a stable and sustainable fixed exchange rate system then becomes an important economic policy. There is an immense literature (for example: Krugman, 1979; Flood and Garber, 1984) on the determinants of a stable fixed exchange rate system. However, theory has far outpaced empirics (for some notable exceptions see Jeanne and Masson, 2000; and Bonet et al., 2005). This paper attempts to enrich the literature by examining the sustainability of the currency board arrangements, which is the strongest form of de jure fixed exchange rate, in Argentina and Hong Kong.

There are three reasons for looking into the exchange rate arrangements of these two economies. First, evidence suggests that currency boards positively impact economic performance. Kwan and Lui (1999) argue that currency boards tend to reduce inflation but at the expense of slower output growth. Ghosh et al. (2000) find that currency boards are associated with better inflation performance and higher output growth. Calomiris and Powell (2001) argue that the Argentine currency board fostered the development of the financial system. Second, similar to Schmukler and Serven (2002), the richness of the history of these two economies enhances our understanding of the impacts of domestic and international events on exchange rate devaluation pressures. Finally, these two economies offer longer and more numerous data series than other economies that have had a currency board (such as Bulgaria, Estonia and Lithuania), which is desirable for performing empirical analysis.

Our empirical strategy makes inferences about the government’s credibility of currency boards based on the interest rate differential between the adopting country and the governing country as well as the economic fundamentals in the adopting country. The government’s credibility is defined as the perceived willingness to enforce the currency board, conditional on economic fundamentals. Employing the Markov switching model proposed in Hamilton (1989), we decompose the pressure

---

3 We aim to contribute to the literature on maintaining fixed exchange rate regimes, but we do not aim to participate in the debate over the choice of the exchange rate. For the literature on the choice of exchange rate regime, see Fischer (2001) and references therein.

4 The exchange rate regimes of these three economies were affected by the creation of Euro in 1999 and making empirical analysis problematic.
on maintaining the currency board (as measured by the interest rate differential) into
two factors: economic fundamentals and market expectations about the strength of the
currency board. Specifically, we assume that market expectations are stochastic with
two possible realizations at any point in time: “credible/no devaluation pressure”
(henceforth good state) and “non-credible/high devaluation pressure” (henceforth bad
state). We also assume that market expectations are unobserved by the
ecnomometrician.

Our results indicate self-fulfilling equilibria in both economies. We argue that
economic fundamentals and expectations are key determinants of the sustainability of
currency boards. We estimate the Markov transition matrix between (good and bad)
states for each economy and compute the expected duration of exchange rate regimes
in each state. We find that once the Hong Kong currency board gains credibility, they
are able to stay credible for longer than the Argentine board on average. Finally, we
show that the relative importance of market expectations and economic fundamentals
is different in these two countries. Market expectations play a more important role
in maintaining the currency board in Argentina than in Hong Kong. Economic
fundamentals, including the trade surplus, real exchange rate and inflation rate are
more important for the sustainability of the Hong Kong currency board.

This paper contributes to the literature on the sustainability of fixed exchange
rate regimes. Krugman (1979) and Flood and Garber (1984) argue that economic
fundamentals are the key determinants in maintaining a fixed exchange rate regime.
On the other hand, Obstfeld (1996) argues that expectations play a more important
role. Rivera-Batiz and Sy (2000) argue that the effectiveness of currency boards relies
on their credibility. Following Jeanne and Masson (2000), we nest these two classes
of models in a unified framework to analyze the relevance of these two types of
factors in determining the success of currency board. Bonet et al. (2005) investigate
the sustainability of the Argentine currency board using a Markov switching model).
We do a similar analysis for Argentina. Our main contribution is to extend this type
of analysis to Hong Kong in order to draw more general conclusions and policy
suggestions for ensuring the sustainability of a currency board.

The remainder of paper is organized as follows. Section 2 discusses the
institutional background of Hong Kong and Argentina. Section 3 describes the data and the estimation methodology. Section 4 presents the empirical results. Section 5 concludes.

2 Institutional and Historical Background

Currency board arrangements stipulate that the domestic currency is fully covered by the central bank's holdings of foreign exchange. Among alternative types of fixed exchange rate systems, currency boards are the most rigid form. Consequently, they should provide the highest degree of exchange rate stability. Hong Kong (in 1983), Argentina (in 1991) and Bulgaria (in 1997) adopted currency boards in order to stabilize their currencies and financial markets. Furthermore, Estonia (in 1992) and Lithuania (in 1995) adopted currency boards and issued new currencies upon declaring independence in order to facilitate their integration into the global economy. Table 1 summarizes the years of operation of the currency boards in these countries and the respective governing currencies. Since we focus on the currency boards of Argentina and Hong Kong in this paper, we will provide a brief historical background of these two currency boards in the remaining part of this section.5

<table>
<thead>
<tr>
<th>Home</th>
<th>Governing Currency</th>
<th>Year</th>
</tr>
</thead>
<tbody>
<tr>
<td>HK</td>
<td>US</td>
<td>1983-Present</td>
</tr>
<tr>
<td>Bulgaria</td>
<td>DM/Euro</td>
<td>1997-1999 (DM)</td>
</tr>
<tr>
<td></td>
<td></td>
<td>1999-present (Euro)</td>
</tr>
<tr>
<td></td>
<td></td>
<td>1999-Present (Euro)</td>
</tr>
<tr>
<td>Lithuania</td>
<td>US/Euro</td>
<td>1995-2001 (US)</td>
</tr>
<tr>
<td></td>
<td></td>
<td>2001-Present (Euro)</td>
</tr>
</tbody>
</table>

Argentina adopted its currency board in 1991 and continued its operation until 2001. Prior to the adoption of the currency board system, Argentina experienced very

5 See Latter (2007) and Greenwood (2008) for recent reviews of the currency board of Hong Kong.
high rates of inflation and slow economic growth. Inflation approached 3,000% in 1989 and 2,300% in 1990, and the government ran a fiscal deficit of 7.6% of GDP. To regain the declining value of Argentine Peso, the Argentine government decided to peg the value of the Peso to the US dollar in April 1991 at the rate of 1Peso/1US$ through a currency board arrangement. The motivation for the currency board was to improve the credibility of the government’s disinflation policy. The hope was that the rigidity of a currency board arrangement would provide an immediate boost to the government’s credibility and help to curb the exorbitant inflation. Under the convertibility law, the central bank was required to hold enough reserves in the form of gold, US dollars, or US dollar-denominated bonds to cover (at least) 100% of base money.

Under the currency board, inflation gradually declined to less than 1% after 1996, and real GDP growth averaged about 3% during the 1990s. The peg had been operating successfully for nearly ten years by the middle of 2001. The Argentine crisis in 2001 caused the credibility of the currency board to deteriorate and rekindled interest in understanding the determinants of the sustainability of currency board systems. Moreover, it cast doubt on the currency board arrangement as a suitable fixed exchange rate system in a highly integrated world economy with massive flows of capital.

After the collapse of the silver standard in 1935, Hong Kong adopted its first currency board arrangement. Due in part to its status as a British colony at the time, the governing currency of this arrangement was the Pound Sterling. This regime was in place for nearly forty years, until 1973, interrupted only during Hong Kong’s occupation by Japan from 1941 to 1945. By the early 1970s, the ongoing depreciation of the Pound Sterling motivated Hong Kong to switch the governing currency to the US dollar. However, this switch failed to halt the depreciation of the Hong Kong dollar relative to other currencies, and as a result the Hong Kong government allowed the Hong Kong dollar to float until 1983.

Owing to weak economic fundamentals and the uncertainty created by the Sino-British negotiations over the sovereignty of Hong Kong in 1983, the HK dollar fell from 5.13 HK$/1US$ in 1981 to $9.8 HK$/1US$ in 1983. Private investors were
in doubt about the business environment that would result when sovereignty was
returned, from Britain to China. This induced capital outflows and slow-down in
investment in Hong Kong. The British government decided to fix the value of the
Hong Kong dollar to the US dollar at the rate of 7.78HK$/1 US$ through a currency
board arrangement. Since then, the currency board in Hong Kong has remained
effective up to the present.

In Hong Kong, there was no monetary authority until 1993, and monetary policy
duties were shared by other institutions. In July 1988, “new accounting arrangements”
were established that required Hong Kong and Shanghai Banking Corporation (HSBC)
to keep an amount of reserve deposits equivalent to its net clearing balance in the
Exchange Fund. The Exchange Fund is, in short, the reserve fund held by the Hong
Kong government to ensure the full convertibility between the Hong Kong dollar and
the US dollar. It ensures that there is enough currency to clear the money market in
order to prevent speculative attacks on the Hong Kong dollar. After the establishment
of Hong Kong Monetary Authority (HKMA) in April 1993, the Exchange Fund Bills
and Exchange Fund Notes were issued to enhance open market operations. The
Liquidity Adjustment Facility (LAF) was set up in June 1992, and the HKMA acting
as lender of last resort began to provide short-term liquidity to banks in trouble
through collateral lending. The LAF rate reflects Hong Kong’s monetary stance on
the foreign exchange market.

3 Data and Descriptive Statistics

The data on Argentina and Hong Kong is obtained from the International
Financial Statistics (maintained by the International Monetary Fund). The sample
period for Argentina is 1991-2001 and for Hong Kong, it is 1984-2005, and the data is
quarterly. The primary variable of interest should be, in principle, the probability of
devaluation, π. However, since we do not observe the probability of devaluation in
reality, we use the expected rate of depreciation as an indicator of devaluation
pressure. As in Svensson (1993) and Jeanne and Masson (2000), we use the interest
rate differential between the domestic economy (Argentina or Hong Kong) and the
respective governing currency country (the United States) to measure the expected
rate of depreciation. More specifically, $\pi = i_{\text{domestic}} - i_{\text{US}}$. Following Calvo and Reinhart (2002), we use the money market rate to measure the interest rate. We use 3 Month Money Market Rate for Argentina, 3 Month Hong Kong Interbank Offer Rate (HIBOR) for Hong Kong and Fed fund rate for the US. A positive (negative) value of the interest differential indicates a pressure towards devaluation (appreciation) as suggested by the uncovered interest rate parity condition.

The pressure to devivate for Argentina and Hong Kong are plotted in Figure 1 and Figure 2, respectively. Figure 1 shows that the expectation of devaluation was high at the time of the adoption of Argentina’s currency board in 1991. Interest rate differential was about 200% in the quarter before the currency board was established, implying high devaluation pressure. Then, the pressure fell sharply as the interest rate differential was stable around 0% throughout the 1990s. The currency board effectively reduced devaluation pressure for a full decade after its implementation. Nonetheless, devaluation pressure begins to creep back up beginning in 2001, as evidenced by the growing interest rate differential. By the time Argentina abandoned the currency board in January 2002, the interest rate differential had increased back to 38%.

Figure 1: Interest rate differential between Argentina and the US

Note: Interest rate differential is computed by $i_{\text{ARG}} - i_{\text{US}}$.

Jeanne and Masson (2000) argue that the interest rate differential is a direct estimate of devaluation probability.
The sample period is from January 1991 to December 2005.

Figure 2: Interest rate differential between Hong Kong and the US

Note: Interest rate differential is computed by $i_{HK} - i_{US}$.

The sample period is from January 1983 to December 2005.

In Figure 2, we observe that Hong Kong experienced several episodes of high devaluation pressure, including the onset of the currency board (1983-84), the first Gulf war (1990-92) and the Asian financial crisis (1997). On the other hand, there were periods with pressure to appreciate, such as the 1987 stock market crash (when the 3 Month HIBOR dropped to less than 2%) and a recent period from 2000 to 2003. Furthermore, the volatility of the expected rate of depreciation became smaller after the Asian Financial Crisis. Gerlach (2005) reaches consistent results using daily data from 1998 to 2001, and suggests that this is due to a calmer external environment and the increased credibility of the HKMA.

The economic fundamentals that we consider include the fiscal balance to GDP ratio ($fiscal$), trade balance to GDP ratio ($trade$), real exchange rate ($rer$, expressed by the ratio of CPI in the home country to CPI in the US), unemployment rate ($unemployment$), inflation rate ($inflation$) and real GDP growth rate ($real\_growth$). $Fiscal$ captures the consistency between fiscal policy and the objective of maintaining the currency board. In the case of Hong Kong, as shown in Figure 3a, the average fiscal balance from 1982 to 2005 was about 3% of GDP. This suggests that government finances have been consistent with the currency board arrangement in
Hong Kong.

On the other hand, there was inconsistency between fiscal policy and the currency board arrangement in Argentina. The government deficit was about 1.4% of GDP during the period 1991-2001. Moreover, as shown in Figure 3b, the fiscal balance deteriorated over time, which worsen the credibility of the currency board.
We consider two indicators of the external sector. We use \textit{trade} to capture the overall performance of the external sector, and employ \textit{rer} to measure the competitiveness of domestic production. A high trade surplus implies a net inflow of foreign capital into the domestic economy, which provides foreign reserve to defend the currency peg. Similarly, a positive trade surplus can be achieved by depreciating the real exchange rate.

Figure 4a shows that the real exchange rate of Hong Kong appreciated and the trade surplus deteriorated over the 15 years following the adoption of the currency board. After the Asian Financial Crisis, the real exchange rate depreciated and hence the trade deficit diminished. Figure 4b shows the real exchange rate of Argentina and the ratio of trade balance to GDP over the period 1991-2002. The real exchange rate appreciated during the period 1991 to 1993, and the trade balance deteriorated. After the Mexican Peso crisis, the real exchange rate depreciated, but the trade deficit only diminished slightly, becoming a surplus by 2001.

\textbf{Figure 4a: Trade Surplus and Real Exchange Rate of Hong Kong, 1982-2005}

\begin{figure}[h]
\centering
\includegraphics[width=\textwidth]{figure4a.png}
\caption{Trade Surplus and Real Exchange Rate of Hong Kong, 1982-2005}
\end{figure}

\textbf{Figure 4b: Trade Surplus and Real Exchange Rate of Argentina, 1991-2001}

\begin{figure}[h]
\centering
\includegraphics[width=\textwidth]{figure4b.png}
\caption{Trade Surplus and Real Exchange Rate of Argentina, 1991-2001}
\end{figure}
Finally, we measure internal economic performance using the unemployment rate, inflation rate and real GDP growth rate. The evolution of the first two of these two variables is shown in Figure 5a for Hong Kong and Figure 5b for Argentina. In the early 1980s, Hong Kong experienced a high inflation rate and unemployment rate. After the adoption of the currency board, a negative relationship between the inflation rate and the unemployment rate became more apparent. After the Asian Financial Crisis, the unemployment rate increased and the economy faced a period of deflation. Consistent with our evidence, Genberg and Pauwels (2004a, b) report that Okun’s law and a Phillip’s curve relationship have held in Hong Kong since the establishment of the currency board.\footnote{Okun’s law shows that there is a negative relationship between output gap and unemployment rate and the Phillip’s curve shows that there is a negative relationship between inflation rate and unemployment rate, in brief.}

In the case of Argentina, the inflation rate dropped dramatically from 1991 to 1993, which suggests that the currency board was successful in controlling inflation. However, the unemployment rate increased over the sample period. Bleaney (2004) argues that this was one of the underlying forces of the collapse of the currency board in Argentina.

**Figure 5a: Inflation Rate and Unemployment Rate of Hong Kong, 1982-2005**
4 Empirical Models

We employ the Markov switching model, developed in Hamilton (1989) and applied in Jeanne and Masson (2000), to detect switches between periods with low and high probabilities of devaluation. Let $\pi_t$ denote the expected devaluation at time $t$. For expositional purpose, we use the term expected devaluation for the rest of the paper even though $\pi<0$ indicates an expectation of appreciation. The expected devaluation depends on the unobserved realization of a stochastic process. Let $s_t$ denote the realization at time $t$, which takes value zero (for good state) or one (for bad state). The expected devaluation are $\alpha_0$ and $\alpha_0+\alpha_d$ in the good and bad state,
respectively. Conceptually, the switching parameters need not be restricted to the intercept of the model. We pursue this parsimonious model because of the small sample size afforded by the quarterly data. The model of two states is also suggested in Jeanne (1997), which derives a theoretical framework with two stable dynamic equilibria when multiple equilibria arise. Our model can be written as

$$\pi_t = \alpha_0 + \alpha_d s_t + x_t \beta + \varepsilon_t$$

where $\varepsilon_t$ is identically and independently distributed with mean zero and variance $\sigma^2$ and $x_t = \{fiscal, trade, rer, unemployment, inflation, real growth\}$. We assume that the state variable, $s_t$, follows a first-order 2-state Markov chain, i.e. the economy is in one of the two possible states at each point in time. Denote the transition probability from state $i$ to state $j$ as $Pr[s_{t+1} = j | s_t = i]$ where $i,j = 0,1$. We assume that the transition probabilities between the two states are constants such that $Pr[s_{t+1} = 0 | s_t = 0] = p$ and $Pr[s_{t+1} = 1 | s_t = 1] = q$.

This further implies that the unconditional probabilities of being in state 0 and state 1 at any time are also constant and are given by

$$Pr[S_t = 0] = \frac{1 - q}{2 - p - q}$$

$$Pr[S_t = 1] = \frac{1 - p}{2 - p - q}$$

The model is estimated by maximum likelihood as proposed in Hamilton (1989). We use the likelihood ratio test proposed by Hansen (1992, 1996) to test the two-state switching model against the fundamental model with no state switching. That is, we test the null hypothesis that $\alpha_d = 0$. A likelihood ratio test of this null hypothesis does not have the usual limiting chi-squared distribution, because the parameters $p$ and $q$ are unidentified under the null. Thus, we cannot employ the standard likelihood ratio (LR) test directly. Hansen (1992, 1996)’s test provides an upper bound to the asymptotic distribution of standardized likelihood ratio statistics, even when conventional regularity conditions (such as unidentified parameters) are violated.
Garcia (1998) provides the adjustment of critical values for a set of specific two-state models. We calculate Hansen's test for all models under the null hypothesis of no state-switching, using a four-lag Newey-West correction. The standardized likelihood ratio tests and their corresponding p-values are reported in Table 3.

5 Empirical Results

5.1 Fundamental Model

We first ignore variations in government credibility and estimate the impact of economic fundamentals on devaluation pressure (as indicated by interest rate differential). The parameter estimates and associated standard errors are presented in Table 2. This model performs poorly across the sample countries. In the case of Argentina, it shows that only the coefficient on trade is statistically significant. On the other hand, the coefficients on rer, unemployment and real growth are statistically significant in the case of Hong Kong.

The estimates from the fundamental model suggest that the currency board of Argentina did not experience pressure to readjust its official exchange rate, whereas that of Hong Kong experienced pressure to devalue. The results for Argentina are particularly counter-intuitive, as Argentina failed to maintain its currency board after 2001. In an ex post sense, the collapse of the currency board of Argentina was due to a currency attack, so that there should have been higher devaluation pressure on the currency board. Thus, the constant term in the regression is expected to be positive, which is not the case in Table 2.

One possible explanation for the poor performance of the fundamental model in capturing the devaluation probability is that the model may be mis-specified. Consistent with Obstfeld (1996) and related studies, changes in market expectations of the sustainability of the fixed exchange rate play an important role in determining the pressure to devaluate. In other words, changes in expected devaluation are associated with the perceived credibility of the central bank to maintain the currency board. In order to capture the effect of expectations, we estimate the model under the assumption that there are multiple equilibria.
5.2 Markov Switching Model

Using the Hansen (1992, 1996) standardized LR test, we reject the hypothesis that government credibility does not vary over time. This suggests that the Markov switching model more accurately describes the dynamics of a currency board than the fundamental model. We infer that the currency boards of Argentina and Hong Kong experience self-fulfilling behavior. Expectations are an important factor in determining the credibility of the currency board.

Looking closer into the case of Argentina, the estimation results in Table 2 show that none of the regressors related to economic fundamentals are statistically significant and that only the intercept term in the bad state is statistically significant, with the expected positive sign. There is no expected depreciation in a state of low devaluation pressure; whereas the changes in expectation due to self-fulfilling behavior induce 41% (i.e. $\alpha_0 + \alpha_d$) expected depreciation in a state of high devaluation pressure. Moreover, the implied average duration of low states and high states are 1 quarter and 50 quarters, respectively. This indicates that the currency board in Argentina experienced persistently high devaluation pressure. These results suggest that the currency crisis in Argentina could have been driven by self-fulfilling behavior, which is related to the credibility of the government rather than economic fundamentals. This interpretation is further supported by the fact that currency premium can be substantial even with hard pegs if they are not fully credible (Edwards, 2000).

Within a Markov switching model, it is possible to estimate how likely the currency is to face more or less devaluation pressure over time, as measured by transition probability. Figure 6a shows the probability of the Argentina Peso entering into a bad (high devaluation) state, where the currency board will lack credibility. These estimates are derived from the estimated Markov Switching model. The probability of being in the bad state decreased upon the adoption of the currency board, (although the probability was still higher than one half for the first four years of the currency board). Beginning in 1995, the currency board maintained a high level of credibility for several years (as the probability of being in bad state kept declining since 1995). However, the probability of getting to the bad state increased dramatically since 2001, which is a year before the collapse of the currency board.
Furthermore, the results of Hansen’s LR test show that the currency board of Hong Kong exhibits self-fulfilling behavior, even though the currency board has been in place since 1983. Moreover, the currency board system in Hong Kong faces a pressure of devaluation of 3% (i.e. $\alpha_0$) and 6% (i.e. $\alpha_0+\alpha_d$) in the low and high states of devaluation, respectively. Although Hong Kong experiences a slight devaluation pressure even in the low state, the credibility of Hong Kong’s currency board is not greatly affected by the self-fulfilling behavior. The implied average duration of low state and high state are 12 quarters and 1 quarter, respectively. Comparing these two economies, the lack of credibility in Argentina may have been associated with the weak economic conditions such as lack of fiscal discipline, deteriorating balance of payments and rigidity of labor market regulations.

Figure 6b plots the smoothed regime probabilities for the expected depreciation of Hong Kong. The probability of being in the high devaluation pressure was high during the period of 1984-1990. Since 1990, the credibility of the currency board of Hong Kong became strong and the probability of being in the state of high devaluation is almost non-existent, except during a five-year period after the Asian Financial Crisis. Interestingly, Crosby (2004) reports that, during the period 1974-2000, the output volatility was the lowest during 1989 to 1995. It suggests that the low output volatility did not realized until the currency board build up its credibility.
Among all the regressors, the coefficients on trade, rer and inflation for the switching model for Hong Kong are statistically significant. The signs are reasonable, with real exchange rate depreciation, higher trade balance to GDP ratios and higher inflation rates reducing the pressure of devaluation. In order to ascertain the economic significance of the estimation results, we compare the impact of each variable on the expected depreciation. Since these regressors are of different units, comparing their respective marginal impact by their regression coefficients is misleading. We therefore normalize the estimates to allow for direct comparison as follow. We normalize by multiplying the respective regression coefficient by one standard deviation of the independent variable. The results answer the following question: how much will devaluation pressure change if there is a one standard deviation change in that variable? This approach of normalization views the independent variables as choice variables by the policy maker, and examines the impacts of the policy change on the expected rate of depreciation. Using this normalization, we find that the real exchange rate is the most important driving force of the credibility of the Hong Kong currency board, followed by the inflation rate and the trade balance. For a one percent appreciation in real exchange rate of Hong Kong relative to the US, the expected depreciation falls by 64%. Similarly, a one percent increase in inflation and trade balance-to-GDP ratio increases the expected depreciation by 31% and 16%, respectively.
The estimated regression results allow us to compute the counterfactual expected depreciation as a function of the observed economic fundamentals. In this exercise, we model a good (low devaluation pressure) state throughout the sample periods in order to isolate the effects from the changes in market expectations. In other words, we examine the case where people do not experience the bad state arrive over the sample period. Accordingly, the counterfactual interest rate differential is the fitted value of interest rate differential at the good state (with the coefficient on the Markov switching parameter set to zero) from the Markov Switching Model. In Figure 7a, we plot the actual and counterfactual expected depreciations for Hong Kong over the sample. The actual and counterfactual expected depreciations diverge during the period 1994 to 1999. This suggests that market expectations were a driving force behind devaluation pressures, rather than economic fundamentals. The high expected depreciation during the Asian Financial Crisis is mainly driven by the self-fulfilling nature of market expectations. Furthermore, the actual and counterfactual expected depreciations track one another more closely after 2003, which suggests a re-establishment of confidence in the currency board. In aggregate, this demonstrates that while proper execution of economic fundamentals is important, lack of credibility can often undermine a currency that would otherwise survive.

**Figure 7a: Actual and Counterfactual Expected Depreciations in HK, 1983-2004**

The actual and counterfactual expected depreciations of Argentina are plotted in
Figure 7b. Although the economic fundamentals are not statistically significant in the regression, the predicted depreciation fits the actual expected depreciation quite well over the sample period. There are several periods with high expected depreciation due to market expectations, including the early period of the currency board, the Mexican Peso crisis in December 1994 and the period before the collapse of the currency board. The expected depreciations during these periods are higher than those predicted from the economic fundamentals. This suggests that self-fulfilling behavior tended to drive expected depreciation.

Figure 7b: Actual and Counterfactual Expected Depreciation of Argentina, 1991-2001

6 Conclusion

This paper has examined the determinants of the sustainability of the currency boards of Argentina and Hong Kong. We suggest that that the failure of the Argentine currency board in 2001 was largely due to the lack of credibility in keeping the currency board functioning, which made the currency board highly vulnerable to expectation changes. On the other hand, in the case of Hong Kong, our empirical results suggest that this economy has experienced a credibility problem, but that the effects of expectations changes are not yet strong enough to cause the collapse of the currency board. Although economic fundamentals are strong enough to maintain the currency board, the government should be careful in managing market expectations.
Future research can investigate the sustainability of fixed exchange rates in a broader context to include less extreme forms of fixed exchange rate systems such as a crawling peg.
### Table 2: Parameter estimates

<table>
<thead>
<tr>
<th></th>
<th>Fundamental Model</th>
<th>Markov switch Model</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>ARG</td>
<td>HK</td>
</tr>
<tr>
<td>Constant</td>
<td>-16.5</td>
<td>3.5***</td>
</tr>
<tr>
<td>Good state</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Difference between good and high state</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Fiscal</td>
<td>-0.83</td>
<td>0.03</td>
</tr>
<tr>
<td>trade</td>
<td>8.83*</td>
<td>0.04</td>
</tr>
<tr>
<td>Real exchange rate</td>
<td>37.9</td>
<td>-0.43***</td>
</tr>
<tr>
<td>Unemployment</td>
<td>-1.00</td>
<td>0.11**</td>
</tr>
<tr>
<td>Inflation</td>
<td>-0.03</td>
<td>0.14</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>-------------------------</td>
<td>----</td>
<td>----</td>
</tr>
<tr>
<td>Real GDP growth</td>
<td>-0.36</td>
<td>-0.05*</td>
</tr>
<tr>
<td></td>
<td>(0.29)</td>
<td>(0.03)</td>
</tr>
<tr>
<td>Implied high state</td>
<td>16.9</td>
<td>6.7</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>P</td>
<td>0.03</td>
<td>0.97</td>
</tr>
<tr>
<td></td>
<td>(0.007)</td>
<td>(0.19)</td>
</tr>
<tr>
<td>Q</td>
<td>0.98</td>
<td>0.66</td>
</tr>
<tr>
<td></td>
<td>(0.023)</td>
<td>(0.017)</td>
</tr>
</tbody>
</table>

Note: ARG and HK are Argentina and Hong Kong, respectively.

The specification for fundamental model is \( \pi_t = \alpha_0 + x_t \beta + \epsilon_t \), whereas that of the Markov switching model is \( \pi_t = \alpha_0 + \alpha_d s_t + x_t \beta + \epsilon_t \). The fundamental models with a constant intercept term whereas the Markov switching models with state dependent intercept term. The implied high state is computed by \( \alpha_0 + \alpha_d \).

We denote the estimates at 10% significance level by *, at 5% significance level **, and at 1% significance level ***.
Table 3: Hansen’s standardized LR test

<table>
<thead>
<tr>
<th></th>
<th>ARG</th>
<th>HK</th>
</tr>
</thead>
<tbody>
<tr>
<td>LR</td>
<td>2.86</td>
<td>3.97</td>
</tr>
<tr>
<td>M=0</td>
<td>0.053</td>
<td>0.002</td>
</tr>
<tr>
<td>M=1</td>
<td>0.090</td>
<td>0.011</td>
</tr>
<tr>
<td>M=2</td>
<td>0.100</td>
<td>0.02</td>
</tr>
<tr>
<td>M=3</td>
<td>0.093</td>
<td>0.021</td>
</tr>
<tr>
<td>M=4</td>
<td>0.102</td>
<td>0.03</td>
</tr>
</tbody>
</table>

Note: See Hansen (1996) for details of the test statistic, such as the definition of M.

The row from M=0 to M=4 are p-values of the LR test statistics.
Reference


Data Appendix:

The data that we used in the study is obtained from the IFS, IMF database.

- Exchange Rate in the database is denominated as home currency to US. Line RF…ZF
- Interest rate differential: Line 60B…ZF (subtract by the governing currency countries’ counter part)
- Fiscal deficit: Line 80…ZF divided by Line 99B…ZF
- Trade balance: Line 78ACDFZF multiple by Line RF…ZF and divided by Line 99B…ZF
- Real exchange rate, \( rer \): Line 64…ZF (divided by the governing currency countries’ counter part)
- Inflation rate (\( inflation \)): Line 64…ZF(Converted to yearly inflation rate by subtracting the same quarter from the previous year)
- Unemployment (\( unemployment \)): Line 67R…ZF
- Real GDP_gr: Line 99BVPZF (Converted to yearly real GDP growth by subtracting the same quarter from the previous year)
<table>
<thead>
<tr>
<th>No.</th>
<th>Year</th>
<th>Authors</th>
<th>Title</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>2006</td>
<td>Helmut Siekmann</td>
<td>The Burden of an Ageing Society as a Public Debt (veröffentlicht in: European Public Law 2007 (13/3))</td>
</tr>
<tr>
<td>2</td>
<td>2006</td>
<td>Helmut Siekmann</td>
<td>Die Unabhängigkeit von EZB und Bundesbank nach geltendem Recht und dem Vertrag über eine Verfassung für Europa</td>
</tr>
<tr>
<td>3</td>
<td>2006</td>
<td>Helmut Siekmann</td>
<td>Die Verwendung des Gewinns der Europäischen Zentralbank und der Bundesbank</td>
</tr>
<tr>
<td>4</td>
<td>2006</td>
<td>Reinhard H. Schmidt, Aneta Hryckiewicz</td>
<td>Financial Systems - Importance, Differences and Convergence</td>
</tr>
<tr>
<td>5</td>
<td>2006</td>
<td>Roman Inderst, Holger M. Mueller, Felix Münnich</td>
<td>Financing A Portfolio of Projects</td>
</tr>
<tr>
<td>6</td>
<td>2006</td>
<td>Roman Inderst, Holger M. Mueller</td>
<td>A Lender-Based Theory of Collateral</td>
</tr>
<tr>
<td>7</td>
<td>2006</td>
<td>Joachim Wieland</td>
<td>Staatsverschuldung als Herausforderung für die Finanzverfassung (veröffentlicht in: JZ 2006, S. 751 ff.)</td>
</tr>
<tr>
<td>8</td>
<td>2007</td>
<td>Helmut Siekmann</td>
<td>Der Anspruch auf Herstellung von Transparenz im Hinblick auf die Kosten und Folgekosten der Steinkohlesubventionierung und den Börsengang der RAG AG</td>
</tr>
<tr>
<td>10</td>
<td>2007</td>
<td>Helmut Siekmann</td>
<td>Die Spielbankabgabe und die Beteiligung der Gemeinden an ihrem Aufkommen – zugleich ein Beitrag zu den finanzverfassungsrechtlichen Ansprüchen der Gemeinden (veröffentlicht in: Organisation und Verfahren im sozialen Rechtsstaat, Festschrift für Friedrich E. Schnapp zum ...</td>
</tr>
<tr>
<td>Band</td>
<td>Jahr</td>
<td>Autor*innen</td>
<td>Thema</td>
</tr>
<tr>
<td>------</td>
<td>------</td>
<td>-------------</td>
<td>-------</td>
</tr>
<tr>
<td>11</td>
<td>2007</td>
<td>Symposium am 26.11.2007 in Frankfurt am Main</td>
<td>Neuordnung der föderalen Finanzbeziehungen</td>
</tr>
<tr>
<td>12</td>
<td>2007</td>
<td>Stefan Gerlach, Peter Kugler</td>
<td>Deflation and Relative Prices: Evidence from Japan and Hong Kong</td>
</tr>
<tr>
<td>13</td>
<td>2007</td>
<td>Katrin Assenmacher-Wesche, Stefan Gerlach, Toshitaka Sekine</td>
<td>Monetary Factors and Inflation in Japan</td>
</tr>
<tr>
<td>14</td>
<td>2007</td>
<td>Guntram B. Wolff</td>
<td>Schuldenanstieg und Haftungsausschluss im deutschen Föderalstaat: Zur Rolle des Moral Hazard</td>
</tr>
<tr>
<td>15</td>
<td>2008</td>
<td>Helmut Siekmann</td>
<td>Föderalismuskommission II für eine zukunftsfähige Gestaltung der Finanzsystem nutzen</td>
</tr>
<tr>
<td>17</td>
<td>2008</td>
<td>Helmut Siekmann</td>
<td>Stellungnahme für die öffentliche Anhörung des Haushaltsausschusses zu dem Gesetzentwurf der Fraktion der SPD und Bündnis 90/Die Grünen für ein Gesetz zur Änderung der Hessischen Landeshaltsordnung</td>
</tr>
<tr>
<td>18</td>
<td>2008</td>
<td>Hans Genberg, Cho-Hoi Hui</td>
<td>The credibility of The Link from the perspective of modern financial theory</td>
</tr>
</tbody>
</table>
| 19   | 2009 | Helmut Siekmann | Stellungnahme für die öffentliche Anhörung des Ausschusses für Wirtschaft, Mittelstand und Energie und des Haushalts- und Finanzausschusses des Landtags Nordrhein-Westfalen
Keine Hilfe für Banken ohne einen neuen Ordnungsrahmen für die Finanzmärkte |
| 20   | 2009 | Chun-Yu Ho, Wai-Yip Alex Ho | On the Sustainability of Currency Boards: Evidence from Argentina and Hong Kong |